

Sleep and CKD in Chinese Adults: A Cross-Sectional Study

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Abstract

Background and objectives To assess the association between self-reported sleep duration and quality and odds of having CKD in Chinese adults on the basis of a community study.

Design, setting, participants, & measurements In this cross-sectional study, we included 11,040 Chinese adults who participated in an ongoing prospective study, the Kailuan cohort. Survey questionnaire items addressed insomnia, daytime sleepiness, snoring, and sleep duration during their 2012 interview. Overall sleep quality was evaluated by summarizing these four sleep parameters. Fasting blood samples and single random midstream morning urine samples were collected in 2012 and analyzed for serum creatinine and proteinuria. CKD was defined by eGFR <60 ml/min per 1.73 m² or proteinuria >300 mg/dl. We also examined those at high or very high risk of having CKD, on the basis of the Kidney Disease Improving Global Outcomes recommendations. The association between sleep quality and CKD was assessed using logistic regression model.

Results Worse overall sleep quality was associated with higher likelihood of being high or very high risk for CKD (multiadjusted odds ratio, 2.69; 95% confidence interval, 1.30 to 5.59 comparing two extreme categories; *P* trend <0.01), but not overall CKD (multiadjusted odds ratio, 1.58; 95% confidence interval, 0.89 to 2.80 comparing two extreme categories; *P* trend =0.46), after adjusting for potential confounders. Specifically, individuals with worse sleep quality were more likely to have proteinuria (multiadjusted odds ratio, 1.95; 95% confidence interval, 1.03 to 3.67 comparing two extreme categories; *P* trend =0.02), rather than lower eGFR level (multiadjusted mean eGFR levels were 96.4 and 93.6 ml/min per 1.73 m² in the two extreme sleep categories, respectively; *P* trend =0.13). However, there was no statistically significant association between individual sleep parameters and CKD status.

Conclusions Worse overall sleep quality was associated with higher odds of being high or very high risk for CKD and proteinuria in Chinese adults.

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Introduction

The prevalence of CKD is estimated to be 8%–16% worldwide (1), and approximately 11% in China (2). Individuals with CKD generally have a higher risk of substantial complications, including progression to ESRD (3), cardiovascular disease (4), and mortality (5). In addition to genetic predisposition factors, risk factors for CKD include presence of certain conditions (*e.g.*, obesity, diabetes, and hypertension) and unhealthy lifestyle factors (*e.g.*, smoking and physical inactivity) (6).

Sleep disorders may be a potential risk factor for CKD because they directly result in chronobiologic alterations in the renin-angiotensin-aldosterone system and sympathetic nervous system activation (7). Furthermore, sleep disorders may indirectly contribute to risk of CKD through influences on the aforementioned diseases (8–10). The notion that sleep disorders may be a risk factor for CKD is supported by the observation that 50%–80% of patients undergoing conventional hemodialysis experience certain

symptoms of sleep disorders, such as insomnia or excessive daytime sleepiness (11). However, only a few studies to date have examined the association between sleep and risk of CKD in the general adult population, and results are inconsistent. For example, in some retrospective studies, nonapnea sleep disorder or short sleep duration were associated with a higher risk of CKD or proteinuria (12,13). In contrast, some studies with modest sample sizes (*n*<500) reported that individuals with short sleep duration, poorer sleep quality, or obstructive sleep apnea (OSA) had a higher kidney filtration rate relative to others (14,15). Although these studies did not control for important confounders (*e.g.*, socioeconomic status [12,13,15] and relevant biochemical markers [9,12,13]), were limited by small sample size (14,15), and had inconsistent results, there is suggestion of a potential link between sleep and kidney function.

We thus conducted a large-scale community-based study including >10,000 Chinese adults to comprehensively assess the association between sleep

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parameters (*i.e.*, insomnia, daytime sleepiness, snoring, and sleep duration) and odds of having CKD, on the basis of the combined indices of eGFR and proteinuria (16). The overall sleep quality score represents a combination of different sleep parameters; thus, it may be a more powerful predictor of health outcomes than any single assessment alone. As a secondary outcome, we also examined four individual sleep parameters.

Materials and Methods

Participants

The current cross-sectional analysis was on the basis of a subset of the Kailuan cohort, a multicenter ongoing cohort including 101,510 Chinese adults (81,110 men and 20,400 women) aged 18–98 years, in 2006–2007, living in Tangshan City, China (17). In the Kailuan cohort, all participants underwent baseline questionnaire assessments of lifestyle and health status and clinical and laboratory examinations in 2006–2007; reassessments were conducted every 2 years. All assessments were conducted in 11 hospitals (referred to as centers) responsible for healthcare of the community. In 2012, information on sleep parameters (including insomnia, daytime sleepiness, snoring, and sleep duration) was collected among 12,990 participants (10,725 men and 2265 women) who underwent examination at one center (*i.e.*, the Kailuan General Hospital). Participants with incomplete information on sleep parameters ($n=265$), serum creatinine or proteinuria ($n=596$), or covariates ($n=989$) were excluded, resulting in 11,040 participants (9197 men and 1843 women) in this analysis. Compared with participants included in the analysis, the participants who were not included in this analysis were older (58.3 versus 53.7 years; $P<0.001$), had lower education levels, higher prevalence of major chronic diseases (such as hypertension, diabetes, history of stroke, or cancer), and included a higher proportion of women (21.4% versus 16.7%; $P<0.001$). The study was jointly approved by the Ethics Committee of the Kailuan General Hospital and the Human Subjects Committee at Brigham and Women's Hospital/Harvard Medical School.

Assessment of CKD

Twelve-hour fasting blood samples were collected before 9:00 a.m. at the face-to-face interview in 2012. Serum creatinine was assessed using the sarcosine oxidase assay method (creatinine kit; BioSino Bio-technology and Science Inc., Beijing, China), with a lower limit detection of 22 $\mu\text{mol/L}$ and an upper limit detection of 3000 $\mu\text{mol/L}$ (linear correlation coefficient ≥ 0.99). Within-laboratory intra- and interassay variable coefficients for serum creatinine were $\leq 5\%$ and $\leq 6\%$, respectively. eGFR was computed using serum creatinine, sex, and age, according to the CKD Epidemiology Collaboration equation:

$$\text{eGFR} = 141 \times \text{minutes} (\text{SCr}/\kappa, 1)^\alpha \times \max (\text{SCr}/\kappa, 1)^{-1.209} \\ \times 0.993^{\text{age}} [\times 1.018 \text{ for women}],$$

where SCr is serum creatinine (in milligram per deciliter), κ is 0.7 for women and 0.9 for men, α is -0.329 for women and -0.411 for men, min is the minimum of SCr/κ or 1, and max is the maximum of SCr/κ or 1 (18).

A single random midstream morning urine sample was collected from each participant during their interview. Urine protein concentration was assessed using the dry-chemistry method and standard urinary sediment examination within 2 hours (H12-MA test strips; Changchun Dirui Medical Technology Co., Ltd., Changchun, China) with a lower limit detection of 15 mg/dl. All of the urine samples were measured using a urine analyzer (N-600; Changchun Dirui Medical Technology Co., Ltd.) at the central laboratory of the Kailuan General Hospital. The results of semiquantitative proteinuria were recorded as negative (<15 mg/dl), trace (15–29 mg/dl), 1+ (30–300 mg/dl), 2+ (300–1000 mg/dl), or 3+ (>1000 mg/dl).

CKD Outcome

In this study, CKD was defined and classified according to the status of eGFR and proteinuria. Participants with $\text{eGFR} < 60$ ml/min per 1.73 m^2 or proteinuria $\geq 2+$ (>300 mg/dl) were considered as having CKD in this analysis. We also examined the association between sleep quality and high or very high risk of CKD, as defined by the Kidney Disease Improving Global Outcomes 2012 recommendations (16). Specifically, the eGFR categories were defined as $\text{eGFR} \geq 90$ ml/min per 1.73 m^2 (G1), 60–89 ml/min per 1.73 m^2 (G2), 45–59 ml/min per 1.73 m^2 (G3a), 30–44 ml/min per 1.73 m^2 (G3b), 15–29 ml/min per 1.73 m^2 (G4), and <15 ml/min per 1.73 m^2 (G5); proteinuria categories were defined as negative and trace (<30 mg/dl, A1), 1+ (30–300 mg/dl, A2), and $\geq 2+$ (>300 mg/dl, A3). Participants with G1–2 and A3, or G3a and A2, or G3b and A1 were defined as being high risk for CKD. Participants with G4–5 and A1, or G3b–5 and A2, or G3a–5 and A3 were defined as being very high risk for CKD.

Assessment of Sleep Parameters (Exposure)

Sleep parameters in this study included insomnia, daytime sleepiness, sleep duration, and snoring. All information on sleep parameters was collected *via* questionnaires using verbal query by a trained interviewer at Kailuan General Hospital healthy examination center during their 2012 interview, as detailed below.

Insomnia

Insomnia status of the participants in the past month was assessed using the Chinese version of the Athens Insomnia Scale (AIS) (19). The AIS is a self-report questionnaire including eight items: the first five items estimate sleep procedure (sleep induction, night awakening, awakening in the early morning, total sleep duration, and total quality of sleep) and the last three items assess decreased sense of wellbeing, overall functioning, and daytime sleepiness. Each item was scored from 0 to 3 (0= no event, 1= mild, 2= moderate, and 3= severe). A total AIS score ≥ 6 was considered as insomnia (20). The AIS has been validated in China (20) and demonstrated good test-retest reliability (83%), sensitivity (96%), and specificity (76%) in Chinese adults, relative to the physician diagnosis on the basis of the Diagnostic and Statistical Manual of Mental Disorders, Fourth Edition (20).

Daytime Sleepiness

We used the Chinese version of Epworth Sleepiness Scale (ESS) to measure daytime sleepiness (21,22). The ESS includes eight questions, each scoring 0–3 with increasing scores signifying higher chance of falling asleep while engaged in specific situations of daily life. A total ESS score ≥ 10 was considered as excessive daytime sleepiness (21). In a validation study of the Chinese ESS, in which 31 Chinese bilingual adults completed both Chinese ESS and English ESS, the Chinese ESS demonstrated high internal consistency (Cronbach $\alpha = 0.81$), good test-retest reliability ($\rho = 0.74$), and significant correlation (Spearman $\rho = 0.75$) with English ESS in Chinese adults (22).

Sleep Duration and Snore Status

We collected information on sleep duration and snoring *via* questions. Participants were asked to report usual total hours of actual sleep during the night. In this analysis, we divided participants into four groups (< 6 , 6–7, 7–8, and ≥ 8 h/d) on the basis of their sleep duration; only approximately 1% of participants reported sleep duration of 9+ h/d.

Snoring was defined in this study as self-reported snoring plus self-reported breathing stops (*i.e.*, apneas); breathing stops were explicitly defined as stops in breathing for > 10 seconds before breathing resumed, occurring more than an estimated 50 times per night. The study definition of snoring was provided to participants and participants self-reported frequency of snoring as never/rare, occasional, or frequent.

Overall Sleep Quality Score

To comprehensively evaluate the association between sleep parameters and odds of having CKD, we further calculated an overall sleep quality score by combining all four sleep parameters, with a total range between 0 (best) and 8 (worst), as detailed in Supplemental Table 1. In this study, the overall sleep score was treated as a continuous variable and a categorical variable; a score of > 5 was considered as worst sleep quality, 3–5 as poor sleep quality, and < 3 as best sleep quality.

Assessment of Covariates

Information on covariates that was collected *via* questionnaires in 2012 included education level (primary, middle, or college), income level (< 600 , 600–1000, or > 1000 RMB/mo), occupation (white collar, blue collar, or coalminer), physical activity (never, < 4 , or ≥ 4 times per week), smoking status (never, past smoker, or current smoker), alcohol consumption (never, past drinker, or current drinker), use of medications (*e.g.*, sleep medications, hypoglycemic agents, and antihypertensives), and history of cancer, myocardial infarction (MI), and stroke.

Weight and height were measured by trained field workers during their 2012 interview; body mass index (BMI) was calculated as weight (kilogram)/height (square meters). Systolic and diastolic BPs were measured twice from the seated position using a mercury sphygmomanometer. The average of the two readings was used for the analyses. Hypertension was defined as systolic BP ≥ 140 mmHg or diastolic BP ≥ 90 mmHg, or use of

antihypertensive medications in last 2 weeks irrespective of BP status. Participants with systolic BP between 120 and 129 mmHg or diastolic BP of 80–89 mmHg were considered as having prehypertension.

The blood concentrations of glucose, triglyceride, LDL cholesterol, and HDL cholesterol were measured by an enzymatic method using an autoanalyzer (Hitachi 747; Hitachi, Tokyo, Japan) at the central laboratory of the Kailuan General Hospital. Diabetes was defined as a fasting blood glucose concentration ≥ 7 mmol/L, or active treatment with insulin or any oral hypoglycemic agent. Impaired fasting glucose was defined as a fasting blood glucose concentration of 5.6–6.9 mmol/L.

Statistical Analyses

Statistical analyses were performed with SAS version 9.2 (SAS Institute, Inc., Cary, NC). Logistic regression was used to calculate odds ratios (ORs) and 95% confidence intervals (95% CIs) for the association between overall sleep quality score and individual sleep parameters and odds of having CKD or proteinuria. We adjusted for age, sex, education level, income level, occupation, physical activity, smoking, alcohol consumption, and BMI, diabetes, hypertension, MI, stroke, cancer, use of sleep medications, and plasma concentrations of triglyceride, LDL, and HDL. Because ORs tend to overestimate the effect size, we also calculated prevalence ratio (PR) using SAS procedure PROC GENMOD (23), adjusting for aforementioned covariates. The association between sleep quality and continuous eGFR was examined using SAS procedure PROC GLM, adjusting for potential confounders.

Because major chronic disease (*e.g.*, diabetes, hypertension, MI, stroke, and cancer) and use of sleep medications might have effects on both sleep disorders and odds of having CKD, we conducted another separate sensitivity analysis by excluding participants with these diseases or conditions.

We also explored potential interactions between overall sleep quality score and age, sex, BMI, diabetes status, and between each sleep parameter, in relation to odds of having CKD or being high/very high risk for CKD, by including multiplicative terms in the multivariate logistic regression models.

Results

The characteristics of the cohort on the basis of CKD status in 2012 are shown in Table 1. Approximately 9% participants reported insomnia (AIS score ≥ 6), approximately 2.0% participants reported daytime sleepiness (ESS score ≥ 10), approximately 40% participants reported any snoring, and the mean sleep duration was 7 h/d (range, 4–14 h/d).

We observed that worse overall sleep quality was associated with higher likelihood of being high or very high risk for CKD (multiadjusted OR, 2.69; 95% CI, 1.30 to 5.59 comparing two extreme categories; P trend < 0.01), but not overall CKD (multiadjusted OR, 1.58; 95% CI, 0.89 to 2.80 comparing two extreme categories; P trend = 0.46), after adjusting for potential confounders (Table 2). Participants with insomnia, daytime sleepiness, shorter/prolonged sleep duration, and frequent snoring also tended

Characteristics	eGFR<60 ml/min per 1.73 m ² or Proteinuria >300 mg/dl		P Value
	No	Yes	
N	10,449	591	
Age, yr	53.4±10.9	58.4±11.7	<0.001
Overall sleep quality score			0.01
	<3	7840 (75.0%)	425 (71.9%)
	3–5	2479 (23.7%)	151 (25.6%)
	>5	130 (1.24%)	15 (2.54%)
Athens Insomnia Scale score			0.04
	<6	9372 (89.7%)	514 (87.0%)
	≥6	1077 (10.3%)	77 (13.0%)
The Epworth Sleepiness Scale score			<0.001
	<10	10257 (98.2%)	567 (95.9%)
	≥10	192 (1.80%)	24 (4.10%)
Sleep duration			0.007
	<6 h/d	1342 (12.8%)	76 (12.9%)
	6–7 h/d	2220 (21.3%)	105 (17.8%)
	7–8 h/d	1850 (17.7%)	85 (14.4%)
	≥8 h/d	5037 (48.2%)	325 (54.9%)
Snoring			0.82
	No	6243 (59.8%)	353 (59.7%)
	Occasional	2507 (23.9%)	137 (23.2%)
	Frequent	1699 (16.3%)	101 (17.1%)
Use of sleep medication			0.83
	Yes	1107 (10.6%)	61 (10.3%)
Education level			<0.001
	Primary school	560 (5.40%)	51 (8.60%)
	High school	8909 (85.3%)	511 (86.5%)
	College	980 (9.30%)	29 (4.90%)
Income level, RMB/mo			<0.001
	<600	2352 (22.5%)	146 (24.7%)
	600–1000	7422 (71.0%)	348 (58.9%)
	>1000	675 (6.5%)	97 (16.4%)
Occupation			0.25
	White collar	697 (6.7%)	41 (6.9%)
	Blue collar	5309 (50.8%)	319 (54.0%)
	Coalminer	4443 (42.5%)	231 (39.1%)
Physical activity, >20 min each time			<0.001
	Never	5110 (48.9%)	212 (35.9%)
	<4 times/wk	3766 (36.0%)	280 (47.4%)
	≥4 times/wk	1573 (15.1%)	99 (16.8%)
Smoking status			0.04
	Never	5614 (53.7%)	330 (55.8%)
	Past smoker	637 (6.1%)	48 (8.20%)
	Current smoker	4198 (40.2%)	213 (36.0%)
Alcohol consumption			0.11
	Never	6003 (57.5%)	365 (61.8%)
	Past drinker	3216 (30.8%)	167 (28.3%)
	Current drinker	1230 (11.8%)	59 (9.9%)
Myocardial infarction history			0.69
	Yes	123 (1.18%)	8 (1.35%)
Stroke history			0.003
	Yes	229 (2.19%)	24 (4.06%)
Cancer			0.87
	Yes	21 (0.20%)	1 (0.17%)
Hypertension			<0.001
	No	1753 (16.8%)	28 (4.74%)
	Prehypertension	3869 (37.0%)	165 (27.9%)
	Yes	4827 (46.2%)	398 (67.3%)
Diabetes			0.03
	No	6882 (65.9%)	293 (49.6%)
	Prediabetes	2437 (23.3%)	144 (23.4%)
	Yes	1130 (10.8%)	154 (26.0%)
Body mass index, kg/m ²	25.0±3.36	25.9±3.47	<0.001
LDL cholesterol, mmol/L	2.10±1.09	2.41±1.13	<0.001
HDL cholesterol, mmol/L	1.41±0.68	1.40±0.65	0.59
Triglyceride, mmol/L	1.73±1.07	2.26±2.05	<0.001

Continuous variables were reported as means±SD. Categorical variables were reported as N (%).

to have higher likelihood of being high or very high risk for CKD. However, the associations were not significant in the fully adjusted models (Table 2). We obtained similar results when PRs were used to estimate the association between sleep quality and CKD status. The multivariate-adjusted PR, comparing two extreme sleep quality categories, was 1.92 (95% CI, 1.08 to 3.42; *P* trend =0.03) for likelihood of being

high or very high risk for CKD and 0.97 (95% CI, 0.66 for 1.43; *P* trend =0.88) for likelihood of CKD, after adjusting for potential confounders (Supplemental Table 2). Specifically, individuals with worse sleep quality were more likely to have proteinuria (multiadjusted OR, 1.95; 95% CI, 1.03 to 3.67 comparing two extreme sleep categories; *P* trend =0.02; Table 3), rather than lower eGFR level

Table 2. The odds ratios (ORs) and 95% confidence intervals (95% CIs) of odds of having CKD, according to sleep status

Sleep parameters	High or Very High Risk for CKD			eGFR<60 ml/min per 1.73 m ² or Proteinuria >300 mg/dl		
	Case/N	Age-Sex Adjusted		Case/N	Age-Sex Adjusted	
		OR (95% CI)	Fully Adjusted ^a OR (95% CI)		OR (95% CI)	Fully Adjusted ^a OR (95% CI)
Overall sleep quality score						
	<3	146/8265	Ref 1	425/8265	Ref 1	Ref 1
	3-5	69/2630	1.46 (1.09 to 1.95)	151/2630	1.07 (0.88 to 1.29)	1.03 (0.82 to 1.22)
	>5	9/145	3.37 (1.68 to 6.78)	15/145	1.88 (1.09 to 3.25)	1.58 (0.89 to 2.80)
	<i>P</i> trend		<0.001		0.12	0.46
AIS score						
	<6	190/9886	Ref 1	514/9886	Ref 1	Ref 1
	≥6	34/1154	1.47 (1.01 to 2.13)	77/1154	1.22 (0.95 to 1.56)	1.18 (0.91 to 1.53)
	<i>P</i> difference		0.04		0.12	0.22
ESS score						
	<10	215/10,824	Ref 1	567/10,824	Ref 1	Ref 1
	≥10	9/216	2.03 (1.03 to 4.02)	24/216	2.08 (1.34 to 3.22)	1.48 (0.93 to 2.35)
	<i>P</i> difference		0.04		0.001	0.10
Sleep duration						
	<6 h/d	39/1418	1.60 (0.99 to 2.59)	76/1418	1.09 (0.79 to 1.49)	1.16 (0.84 to 1.61)
	6-7 h/d	40/2325	1.05 (0.65 to 1.68)	105/2325	0.98 (0.73 to 1.32)	1.06 (0.78 to 1.43)
	7-8 h/d	31/1935	Ref 1	85/1935	Ref 1	Ref 1
	≥8 h/d	114/5362	1.22 (0.82 to 1.83)	325/5362	1.25 (0.97 to 1.59)	1.19 (0.92 to 1.54)
	<i>P</i> trend		0.78		0.03	0.21
Snoring						
	No	123/6596	Ref 1	353/6596	Ref 1	Ref 1
	Occasional	52/2644	1.12 (0.81 to 1.56)	137/2644	1.36 (0.85 to 1.27)	0.98 (0.79 to 1.22)
	Frequent	49/1800	1.50 (1.07 to 2.11)	101/1800	1.05 (0.83 to 1.31)	0.95 (0.75 to 1.22)
	<i>P</i> trend		0.02		0.66	0.71

AIS, Athens Insomnia Scale; ESS, the Epworth Sleepiness Scale.

^aAdjusted for age, sex, education level (primary, high school, or college), income level (<600, 600-1000, or >1000 RMB/mo), occupation (white collar, blue collar, or coalminer), physical activity (never, <4, or ≥4 times/week), smoking status (never, past smoker, or current smoker), alcohol consumption (never, past drinker, or current drinker), myocardial infarction history (no, yes), stroke history (no, yes), cancer history (no, yes), hypertension (no, prehypertension, or hypertension), diabetes (no, prediabetes, or diabetes), use of sleep medication (no, yes), body mass index (<24, 24-28, or ≥28 kg/m²), triglyceride (<0.82, 0.82-1.22, 1.22-1.87, or ≥1.87 mmol/L), LDL (<1.76, 1.76-2.12, 2.12-2.65, or ≥2.65 mmol/L), and HDL (<1.31, 1.31-1.54, 1.54-1.79, or ≥1.79 mmol/L). Each set of sleep variables was modeled separately with adjustment for the covariates.

Overall Sleep Quality Score	Case/N	Multivariate Adjusted Model 1 ^a	Multivariate Adjusted Model 2 ^b
<3	240/8265	Ref 1	Ref 1
3–5	109/2630	1.21 (0.95 to 1.54)	1.21 (0.95 to 1.54)
>5	12/145	1.95 (1.03 to 3.67)	1.81 (0.95 to 3.44)
<i>P</i> trend		0.02	0.03

^aAdjusted for age, sex, education level (primary, high school, or college), income level (<600, 600–1000, or >1000 RMB/mo), occupation (white collar, blue collar, or coalminer), physical activity (never, <4, or ≥4 times/week), smoking status (never, past smoker, or current smoker), alcohol consumption (never, past drinker, or current drinker), myocardial infarction history (no, yes), stroke history (no, yes), cancer history (no, yes), hypertension (no, prehypertension, or hypertension), diabetes (no, prediabetes, or diabetes), use of sleep medication (no, yes), body mass index (<24, 24–28, or ≥28 kg/m²), triglyceride (<0.82, 0.82–1.22, 1.22–1.87, or ≥1.87 mmol/L), LDL (<1.76, 1.76–2.12, 2.12–2.65, or ≥2.65 mmol/L), and HDL (<1.31, 1.31–1.54, 1.54–1.79, or ≥1.79 mmol/L).

^bFurther adjusted for eGFR (ml/min per 1.73 m²).

(adjusted mean eGFR levels were 96.4 and 93.6 ml/min per 1.73 m² in two extreme sleep categories, respectively; *P* trend =0.13; Table 4).

In the sensitivity analyses, exclusion of the participants with diabetes, hypertension, using sleep medications, MI, stroke, and cancer generated similar results (Supplemental Tables 3–5). We did not find significant interactions between age, sex, BMI, diabetes, and the overall sleep quality score, and between each sleep parameters, in relation to odds of having CKD or being high or very high risk for CKD (*P* interaction >0.05 for all).

Discussion

In this large-scale community-based study including >10,000 participants, we found that higher overall sleep quality score (*i.e.*, poorer sleep quality) was associated with higher likelihood of being high or very high risk for CKD. The association appeared to be driven by proteinuria rather than lower eGFR. Likewise, participants with insomnia, daytime sleepiness, shorter/prolonged sleep duration, and frequent snoring also tended to have higher likelihood of CKD. However, the associations were not significant.

Our findings are consistent with previous studies that individuals with CKD had a higher prevalence of shorter

and higher fragmented sleep and excessive daytime sleepiness, relative to those without CKD (24). In a recent meta-analysis including six observational studies, short sleepers tended to have higher risk of having CKD (pooled OR, 1.51; 95% CI, 0.99 to 2.55) (25). Excessive daytime sleepiness and short sleep duration were also found to be associated with greater 24-hour urinary albumin excretion (26). Similarly, a small sample study (*n*=91) linked snoring to proteinuria and high serum urea and creatinine concentration (27). High prevalence of OSA was found in non-dialysis patients with CKD and the prevalence and severity of OSA increased with progression of CKD stage (28,29). It was also observed that severe OSA was also independently associated with higher serum cystatin C levels in patients without CKD; serum cystatin C is considered to be a biomarker reflecting clinically latent renal dysfunction (30).

Interestingly, several previous studies reported that objective shorter sleep duration or reported apnea-hypopnea index and desaturation index were associated with higher kidney filtration rates (14), or higher eGFR and urine albumin-to-creatinine ratio (15). Previous studies suggested that high levels of eGFR might be an early predictor of diabetes and hypertension (7,10), and were associated with higher incidence of coronary heart disease and mortality

Overall Sleep Quality Score	<i>N</i>	Multivariate Adjusted Model 1 ^a	Multivariate Adjusted Model 2 ^b
<3	8265	96.4 (92.3 to 100.6)	89.4 (85.1 to 93.8)
3–5	2630	96.5 (92.4 to 100.6)	89.6 (85.2 to 93.9)
>5	145	93.6 (88.6 to 98.5) ^c	87.3 (82.2 to 93.4)
<i>P</i> trend		0.13	0.30

^aAdjusted for age, sex, education level (primary, high school, or college), income level (<600, 600–1000, or >1000 RMB/mo), occupation (white collar, blue collar, or coalminer), physical activity (never, <4, or ≥4 times/week), smoking status (never, past smoker, or current smoker), alcohol consumption (never, past drinker, or current drinker), myocardial infarction history (no, yes), stroke history (no, yes), cancer history (no, yes), hypertension (no, prehypertension, or hypertension), diabetes (no, prediabetes, or diabetes), use of sleep medication (no, yes), body mass index (<24, 24–28, or ≥28 kg/m²), triglyceride (<0.82, 0.82–1.22, 1.22–1.87, or ≥1.87 mmol/L), LDL (<1.76, 1.76–2.12, 2.12–2.65, or ≥2.65 mmol/L), and HDL (<1.31, 1.31–1.54, 1.54–1.79, or ≥1.79 mmol/L).

^bFurther adjusted for presence of proteinuria (yes/no).

^c*P*<0.05, relative to overall sleep quality score <3.

(31). In this context, sleep disorders might increase eGFR to abnormally high levels, and subsequently increase the risk for glomerular hypertension, renal injury, and the development of CKD (14). We thus performed sensitivity analyses by excluding participants with diabetes and hypertension, and observed similar relationships between overall sleep quality score and CKD. These findings, together with previous studies, suggest that sleep disorders could damage the renal function in a stepwise manner over a long-term period.

The exact underlying mechanism of association between sleep and CKD remains unknown. There were some lines of possible evidence in support of adverse effect of sleep on development of CKD. As described above, glomerular hyperfiltration, a mark of early kidney damage and risk factor for proteinuria, is observed in individuals with sleep disorders (*e.g.*, OSA) (14,15). Sleep disorders abnormally activate the autonomic nervous system, a key regulator of kidney function, and then produce unfavorable effects on renal hemodynamics and BP (32). It has been observed that individuals with sleep disorders have excessively elevated daytime sympathetic nervous system activity (33) and fail to increase the cardiac vagal tone during the transition from wakefulness to nonrapid eye movement (34). Furthermore, snoring, a symptom of sleep-related breathing disorder, has a notable adverse effect on the regulation of renin-angiotensin-aldosterone system through pathways of hypoxemia, oxidative stress, and activated sympathetic activities, which result in endothelial dysfunction and lead to renal damage (10). Finally, short/prolonged sleep duration and sleep disorders could lead to systemic inflammation (35,36), which potentially results in glomerular endothelial dysfunction and subsequent renal dysfunction (37).

Our study had several limitations. First, this study was cross-sectional, which does not allow for establishing causality of the observed associations. The association between sleep and CKD might be bidirectional. Second, all sleep information was collected *via* self-report, and was subject to recall bias and misclassification, particularly for sleep duration and OSA (we used snoring as a surrogate). Third, the participants who were not included in this analysis were older and had high prevalence of major chronic diseases, which were associated with a higher risk of CKD. Therefore, we might underestimate the association between sleep and CKD. Fourth, the detection of proteinuria was on the basis of a semiquantitative method from single random urine samples rather than continuous quantitative method using 24-hour urine sample, which could lead to misclassification of CKD. Because of the high sensitivity (93.3%) and specificity (91.6%) of the semiquantitative method relative to the quantitative method (38), the approach used in this study remains a useful method for detecting proteinuria, especially in an epidemiologic study with large sample size.

In conclusion, in this large-scale community-based study we found that participants with lower overall sleep quality had a high likelihood of being high or very high risk for CKD. Future prospective studies are warranted to clarify the temporal relationship between objectively measured sleep disorders and CKD. If our findings are confirmed, sleep disorders as modifiable lifestyle factors may be a novel intervention target for preventing CKD.

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Disclosures

None.

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