Diet Soda Consumption and Risk of Incident End Stage Renal Disease

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Lydia A. Bazzano,* Josef Coresh,*‡** and Lawrence J. Appel*‡**

Abstract

Background and objectives Diet soda consumption is common in the United States and is associated with impaired glucose metabolism, diabetes, and metabolic syndrome.

Design, setting, participants, & measurements We prospectively analyzed diet soda consumption, assessed by food frequency questionnaire at baseline (1987–1989) and a follow-up examination (1993–1995), and incident ESRD through December 31, 2012 in the Atherosclerosis Risk in Communities study (n=15,368).

Results Baseline mean age of participants was 54 years, 55% were female, and 27% were black. The majority of participants (43.5%) consumed <1 glass/wk of diet soda; 17.8% consumed 1–4 glasses/wk; 25.3% consumed 5–7 glasses/wk; and 13.5% consumed >7 glasses/wk. Over a median follow-up of 23 years, 357 incident ESRD cases were observed. Relative to <1 glass/wk of diet soda, consuming 1–4 glasses/wk, 5–7 glasses/wk, and >7 glasses/wk, respectively, was associated with 1.08-times (95% CI, 1.01 to 1.75), 1.33-times (95% CI, 1.01 to 1.75), and 1.83-times (95% CI, 1.01 to 2.52) higher risk of ESRD after adjusting for age, sex, race-center, education level, smoking status, physical activity, total caloric intake, eGFR, body mass index category, diabetes, systolic BP, and serum uric acid (P value for trend <0.001). Results were similar after additional adjustment for dietary intake of added sugar, diet quality, dietary sodium, dietary fructose, sugar-sweetened beverages, and dietary phosphorus. Risk estimates were similar by body mass index category (P value for interaction = 0.82), but the association between diet soda and ESRD was only significant for those who were overweight or obese at baseline. Sugar-sweetened beverage consumption was not significantly associated with ESRD in the fully adjusted model.

Conclusions Diet soda consumption was associated with higher ESRD risk in this general population sample. Further research is necessary to validate these findings in other study populations and to examine potential mechanisms through which diet soda could impact kidney disease.


Introduction

Soft drinks are a major source of calories (5% of total caloric intake) and the primary source of added sugar (33% of total added sugar intake) in the United States diet (1). These types of beverages are associated with weight gain, type 2 diabetes mellitus, and cardiovascular disease (2–4). Therefore, the 2015 Dietary Guidelines for Americans and the American Heart Association recommend limiting dietary intake of added sugars, in part, by avoiding soft drinks and other sugar-sweetened beverages (5,6). Numerous policy initiatives have been implemented (e.g., taxing the purchase of soda) to reduce sugar-sweetened beverage consumption at the population-level and to fund health promotion programs (7,8).

As a result of the widespread public awareness of the sugar content, caloric burden, and adverse health consequences of regular soda, diet soda has become an increasingly common substitute in the United States (8,9). Consumption of diet beverages is higher in North America than any other region in the world (8). Recent studies have shown that diet soda and artificial sweeteners contained in diet soda may adversely affect glucose levels, and may increase the risk of developing metabolic syndrome and diabetes, in part, through impairment of glucose and increase in waist circumference (10–12). With respect to kidney disease, the results of the few existing studies on diet soda have been inconsistent and the overall evidence base is inconclusive (13–16).

The objective of this study was to investigate the relationship between diet soda with the development of incident ESRD in a general population sample and to assess the independence of this association from established ESRD risk factors and dietary factors related to diet soda.

Materials and Methods

Study Design

The Atherosclerosis Risk in Communities (ARIC) study is a prospective cohort of 15,792 middle-aged...
(45–64 years), predominantly black and white men and women from four United States communities: Forsyth County, North Carolina; Jackson, Mississippi; suburbs of Minneapolis, Minnesota; and Washington County, Maryland (17). Participants enrolled in 1987–1989 (baseline, visit 1) and follow-up visits occurred in 1990–1992 (visit 2), 1993–1995 (visit 3), and 1996–1998 (visit 4), with the most recent visit in 2011–2013 (visit 5). The protocol was approved by the Institutional Review Board. Procedures were followed in accordance with the Declaration of Helsinki.

Study Population
After excluding participants with missing diet data (≥10 food items not reported) or implausible total caloric intake (women: <500 or >3500 kcal; men: <700 or >4500 kcal) (n=364); those who were neither black nor white (n=47); and those with baseline eGFR <45 ml/min per 1.73 m² (n=13), the analytic sample size was 15,368.

Assessment of Diet Soda Intake
Usual dietary intake was assessed by a semiquantitative, 66-item food frequency questionnaire (FFQ) administered by trained interviewers at baseline (1987–1989, visit 1) and a follow-up visit (1993–1995, visit 3) (18). Participants reported how often they consumed food items on average over the past year. Visual aids (glasses, measuring cups) were used to illustrate portion sizes in order to improve accuracy. The reliability of this FFQ was previously demonstrated in a random subset of 419 ARIC study participants who repeated the FFQ at a follow-up visit (visit 2) (19). The cumulative average diet, incorporating data from both assessments of dietary intake (baseline and visit 3), was used to depict beverage intake (20).

Diet soda was described on the FFQ as one 8-ounce glass of low-calorie soft drinks such as Diet Coke, Diet Pepsi, or Diet 7-Up. To provide a more complete assessment of beverage intake, we also assessed sugar-sweetened beverage consumption, which consisted of regular soft drinks (Coke, Pepsi, 7-Up, or ginger ale) as well as fruit-flavored or noncarbonated beverages (lemonade, Kool-Aid, or Hawaiian Punch). Consumption frequency was categorized as <1 glass/wk, 1–4 glasses/wk, 5–7 glasses/wk, and >7 glasses/wk.

Ascertainment of Incident ESRD
Incident ESRD was defined as the initiation of RRT (transplant, dialysis) between baseline (1987–1989) and December 31, 2012 as identified by linkage with the US Renal Data System (USRDS) registry. Study participants were censored at the date of death or the end of the observation period for this study (December 31, 2012). In a validation study of this definition of ESRD compared with physician-determined treated kidney failure on the basis of medical chart review in the ARIC study, sensitivity was 95% and specificity was 100% (21).

As a secondary outcome, incident CKD was defined as meeting one of the following criteria: (1) eGFR <60 ml/min per 1.73 m² at any follow-up visit accompanied by ≥25% eGFR decline, (2) CKD-related hospitalization using International Classification of Diseases-9/10 codes, (3) CKD-related death using International Classification of Diseases-9/10 codes, or (4) USRDS-identified ESRD (22).

Measurement of Covariates
Demographic characteristics (age, sex, race), socioeconomic status (education), health behaviors (smoking, physical activity), and health history (disease diagnosis, medication use) were captured using a structured questionnaire administered by trained interviewers at baseline. A modified Baecke questionnaire was used to create an index of leisure-time sports and exercise, incorporating frequency, duration, and intensity of each type of activity on average over the preceding year (23).

We quantified diet quality with the Alternative Healthy Eating Index 2010, which we modified to exclude alcohol given its missingness (45%) (24). Higher scores represent higher diet quality. Dietary intake of protein, phosphorus, potassium, magnesium, calcium, sodium, and fructose was calculated by combining frequency of consumption, portion size, and nutritional content of each food item on the FFQ. Dietary acid load was estimated with the Remer and Manz equation for potential renal acid load: 0.49 × protein + 0.037 × phosphorus − 0.021 × potassium − 0.026 × magnesium − 0.013 × calcium (25–28).

Weight and height were measured while participants wore light clothing and no shoes, and body mass index (BMI) was categorized as normal (<25 kg/m²), overweight (25 to <30 kg/m²), or obese (≥30 kg/m²). After resting for at least 5 minutes, three measurements of BP were taken using a random-zero sphygmomanometer by a certified technician and the average of the second and third readings was used. Hypertension was defined as systolic BP ≥140 mmHg, diastolic BP ≥90 mmHg, or current use of antihypertensive medication in the preceding 2 weeks. Serum glucose was quantified by the modified hexokinase/glucose-6-phosphate dehydrogenase method. Diabetes was defined as fasting blood glucose ≥126 mg/dl, nonfasting blood glucose ≥200 mg/dl, self-report of diagnosed diabetes, or current use of diabetes medication in the preceding 2 weeks. Serum creatinine was measured by the modified kinetic Jaffe method and standardized to the National Institute of Standards and Technology standard (29). eGFR was estimated using the 2009 CKD Epidemiology Collaboration equation on the basis of creatinine (30). Serum uric acid was measured using the uricase method (31).

Statistical Analyses
Baseline characteristics were examined according to categories of diet soda consumption using descriptive statistics, and differences were tested using a nonparametric test for trend across ordered groups (32).

We used Cox proportional hazards regression to calculate hazard ratios (HR) and 95% confidence intervals (95% CI) for the association between diet soda consumption and incident ESRD risk, incorporating time until ESRD. We conducted a test of linear trend using the median value of diet soda consumption within each category in the regression models. In addition to categorical analysis, diet soda intake was modeled continuously and effect estimates were expressed per one additional glass consumed. We conducted stratified analyses and tested for interaction by sex, race, diabetes status, and BMI category.

Five successive regression models were constructed. Model 1 included demographic characteristics (age, sex, race-center), education level as a proxy for socioeconomic
Table 1. Baseline characteristics according to categories of diet soda consumption

<table>
<thead>
<tr>
<th>Characteristics</th>
<th>Categories of Diet Soda Consumption, glasses/wk</th>
<th></th>
<th></th>
<th></th>
<th>P Value for Trend</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td>&lt;1 (n=6678; 43.5%)</td>
<td>1–4 (n=2728; 17.8%)</td>
<td>5–7 (n=3885; 25.3%)</td>
<td>&gt;7 (n=2077; 13.5%)</td>
<td></td>
</tr>
<tr>
<td>Age, yr</td>
<td>54.4 (5.8)</td>
<td>54.8 (5.6)</td>
<td>54.0 (5.7)</td>
<td>52.9 (5.6)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Female, % (n)</td>
<td>52.2 (3486)</td>
<td>55.5 (1515)</td>
<td>58.4 (2268)</td>
<td>58.3 (1210)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Black, % (n)</td>
<td>34.6 (2311)</td>
<td>21.6 (588)</td>
<td>23.7 (921)</td>
<td>14.1 (292)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>High school education, % (n)</td>
<td>72.2 (4810)</td>
<td>81.3 (2215)</td>
<td>78.3 (3040)</td>
<td>79.6 (1651)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Current smoker, % (n)</td>
<td>33.3 (2222)</td>
<td>19.0 (517)</td>
<td>19.4 (754)</td>
<td>24.5 (508)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Physical activity index</td>
<td>2.37 (0.78)</td>
<td>2.49 (0.80)</td>
<td>2.49 (0.81)</td>
<td>2.43 (0.79)</td>
<td>0.001</td>
</tr>
<tr>
<td>Fasting glucose, mg/dl</td>
<td>102.8 (26.7)</td>
<td>104.9 (29.4)</td>
<td>111.1 (42.6)</td>
<td>114.3 (42.0)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Diabetes, % (n)</td>
<td>7.7 (505)</td>
<td>8.9 (242)</td>
<td>16.1 (621)</td>
<td>20.8 (430)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>SBP, mmHg</td>
<td>122.0 (19.6)</td>
<td>120.5 (17.6)</td>
<td>121.2 (18.8)</td>
<td>119.9 (17.7)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Hypertension, % (n)</td>
<td>34.8 (2315)</td>
<td>33.4 (905)</td>
<td>36.6 (1,414)</td>
<td>34.6 (717)</td>
<td>0.65</td>
</tr>
<tr>
<td>BMI, kg/m²</td>
<td>26.8 (5.2)</td>
<td>27.5 (5.0)</td>
<td>28.4 (5.4)</td>
<td>29.4 (5.7)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>BMI category</td>
<td></td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Normal (&lt;25 kg/m²), % (n)</td>
<td>40.0 (2655)</td>
<td>34.1 (931)</td>
<td>27.1 (1051)</td>
<td>21.2 (440)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Overweight (25 to &lt;30 kg/m²), % (n)</td>
<td>38.1 (2543)</td>
<td>40.4 (1101)</td>
<td>40.7 (1581)</td>
<td>40.1 (832)</td>
<td></td>
</tr>
<tr>
<td>Obese (≥30 kg/m²), % (n)</td>
<td>22.1 (1477)</td>
<td>25.5 (696)</td>
<td>32.2 (1249)</td>
<td>38.6 (801)</td>
<td></td>
</tr>
<tr>
<td>eGFR, ml/min per 1.73 m²</td>
<td>103.2 (16.1)</td>
<td>101.6 (14.1)</td>
<td>102.1 (15.7)</td>
<td>102.4 (15.0)</td>
<td>0.21</td>
</tr>
<tr>
<td>Dietary acid load, mEq/d</td>
<td>3.7 (12.6)</td>
<td>3.2 (11.5)</td>
<td>5.0 (12.1)</td>
<td>8.1 (13.6)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Diet quality score</td>
<td>42.4 (9.7)</td>
<td>45.4 (9.2)</td>
<td>45.6 (9.1)</td>
<td>44.6 (9.2)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>SSB, glasses/wk</td>
<td>5.3 (6.9)</td>
<td>3.0 (4.2)</td>
<td>2.5 (4.0)</td>
<td>2.9 (5.6)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Dietary phosphorus, mg/d</td>
<td>1021 (392)</td>
<td>1070 (372)</td>
<td>1088 (370)</td>
<td>1214 (407)</td>
<td>&lt;0.001</td>
</tr>
<tr>
<td>Total caloric intake, kcal</td>
<td>1645 (572)</td>
<td>1584 (525)</td>
<td>1565 (521)</td>
<td>1680 (566)</td>
<td>0.02</td>
</tr>
</tbody>
</table>

Data are displayed as mean (SD) for continuous variables or indicated as % (n) for categorical variables. SBP, systolic BP; BMI, body mass index; mEq, milliequivalents; SSB, sugar-sweetened beverage.
status, health behaviors (smoking status, physical activity), total caloric intake as the standard method for energy adjustment, and baseline eGFR modeled as two linear spline terms with one knot at 90 ml/min per 1.73 m² considering the nonlinear relationship between eGFR and kidney disease risk and the relatively high level of kidney function in this general population sample (20,33). The race-center interaction term was used given the nonuniform distribution of racial groups across study sites. Model 2 additionally adjusted for comorbidities (BMI category, diabetes, systolic BP, serum uric acid). Models 3a–c investigated the independence of the association between diet soda consumption and ESRD after accounting for dietary factors related to diet soda. To avoid collinearity between multiple dietary factors, they were added to the regression model separately. Model 3a adjusted for dietary acid load in addition to Model 2 covariates. Model 3b adjusted for diet quality (modified Alternative Healthy Eating Index 2010), dietary intake of sodium, dietary intake of fructose, and frequency of sugar-sweetened beverage consumption in addition to Model 2 covariates. Model 3c adjusted for dietary intake of phosphorus and the Model 2 covariates. Analyses were performed using Stata statistical software version 14.1 (StataCorp LP, College Station, TX).

Results
In the overall study population, baseline mean age was 54 years, 55% were female, 27% were black, 12% had diabetes, 35% had hypertension, and baseline mean eGFR was 102.5 ml/min per 1.73 m². The majority of participants (43.5%) consumed <1 glass/wk of diet soda, 17.8% consumed 1–4 glasses/wk, 25.3% consumed 5–7 glasses/wk, and 13.5% consumed >7 glasses/wk (Table 1). Those who consumed the highest amount of diet soda (>7 glasses/wk) were more likely to be female, white, and obese, and to have diabetes. Higher frequency of diet soda consumption was associated with lower intake of sugar-sweetened beverages and higher dietary intake of phosphorus and dietary acid load (P < 0.001).

Over a median follow-up of 23 years, there were 357 incident ESRD cases. Relative to <1 glass/wk of diet soda, consuming 1–4 glasses/wk, 5–7 glasses/wk, and >7 glasses/wk, respectively, was associated with 1.88-times (95% CI, 1.01 to 2.88) higher risk of ESRD after adjusting for age, sex, race-center, education level, smoking status, physical activity, total caloric intake, eGFR, BMI category, diabetes, systolic BP, and serum uric acid (Model 2, P value for trend < 0.001; Table 2). In the continuous analysis, for each additional glass of diet soda consumed per day, there was a 29% higher risk of ESRD (Model 2, HR, 1.29; 95% CI, 1.16 to 1.43; P < 0.001). In both categorical and continuous analyses, effect estimates were similar after additional adjustment for dietary acid load (Model 3a), after accounting for diet quality, dietary sodium, dietary fructose, and sugar-sweetened beverage consumption (Model 3b), and after adjusting for dietary intake of phosphorus (Model 3c). Results were similar, although attenuated, for the association between diet soda consumption and the secondary outcome of incident CKD (Supplemental Table 1).
ESRD risk for the highest versus lowest categories of diet soda consumption were similar by sex (P value for interaction = 0.59) and racial group (P value for interaction = 0.41; Figure 1). Associations were slightly stronger among those with diabetes and not significant among those without diabetes, but there was no statistical evidence of interaction (P value for interaction = 0.34). Although risk estimates were similar by BMI category (P value for interaction = 0.82), the association between diet soda and ESRD was only significant for those who were overweight or obese at baseline.

In the minimally adjusted model (Model 1), the highest (>7 glasses/wk) versus lowest (<1 glass/wk) consumption category for sugar-sweetened beverages appeared to be associated with lower risk of incident ESRD (Table 3). This association was no longer statistically significant in nearly all subsequent models with further adjustment.

Discussion
In this diverse, community-based population of 15,368 black and white men and women, higher consumption of diet soda was associated with a graded risk of developing ESRD over a median follow-up of 23 years. This dose-response relationship was independent of several known ESRD risk factors and other dietary factors related to diet soda, and results were consistent in population subgroups.

To the best of our knowledge, this study is the first to report an association between diet soda and incident ESRD. There is a paucity of literature on diet soda and other artificially-sweetened beverages, and only a few studies have related this beverage type to kidney outcomes. In a subset of 3318 participants in the Nurses’ Health Study, a study population which is predominantly white and exclusively female, consuming ≥2 artificially-sweetened beverages per day relative to <1 per month was associated with a 2-fold higher risk of ≥30% eGFR decline after adjusting for age, caloric intake, hypertension, BMI, diabetes, cigarette smoking, physical activity, and cardiovascular disease (odds ratio, 2.02; 95% CI, 1.36 to 3.01) (14). These investigators found a similar association with rapid eGFR decline defined as ≥3 ml/min per 1.73 m² per year (odds ratio, 2.20; 95% CI, 1.36 to 3.55). A case-control study was conducted in 1980–1983 with 465 CKD patients from four North Carolina hospitals and 467 community-dwelling controls with frequency matching on age, sex, race, and proximity to a study hospital (15). After restricting the analysis to self-respondents only (214 cases and 422 controls), drinking ≥2 artificially-sweetened sodas/d relative to never or <1 drink/wk was significantly associated with CKD after adjusting for matching factors, BMI, income, education, analgesic use, and diabetes (odds ratio, 4.21; 95% CI, 1.21 to 14.61). In a cross-sectional analysis of 9358 participants in the 1999–2004 National Health and Nutrition Examination Survey (NHANES), diet soda consumption was not associated with albuminuria after adjusting for sugar-sweetened soda consumption, age, race-ethnicity, gender, and poverty status (odds ratio, 0.94; 95% CI, 0.64 to 1.39) (16). Inconsistency in the literature may be due to differences in study design, outcome definition, exposure classification, and covariates. Replicating these findings will be essential to establishing more definitive knowledge about the kidney health implications of artificially-sweetened beverage intake.

There are several potential mechanisms through which diet soda could cause renal damage. Sodas (both diet and regular) contain phosphorus as an additive for color and flavor (34). Dietary phosphorus may affect serum levels of phosphorus and fibroblast growth factor-23 (35,36). In a separate analysis of the ARIC study, the highest versus lowest quintile of fibroblast growth factor-23 was associated with a 2-fold higher risk of incident ESRD (37). Driven by its phosphorus content, diet soda could increase dietary acid load and thereby increase kidney disease risk (27,28). In another ARIC study analysis, the highest versus lowest quartile of dietary acid load was associated with 1.13-times higher risk of developing CKD (26). Alternatively, high consumption of diet soda could be perceived as a proxy for poor diet quality, considering that diet soda is often consumed as a substitute for sugar-sweetened beverages in an attempt to reduce caloric intake and body weight. Poor diet quality has been assessed by various indices and shown to increase the risk of albuminuria and kidney function decline (38). However, we evaluated these dietary factors in multivariable regression models and found that the association between diet soda and ESRD was independent of diet quality, dietary acid load, and dietary intake of phosphorus.

Diet soda consumption could also plausibly affect kidney disease risk by modifying glucose metabolism. In the ARIC study, higher consumption of diet soda was associated with incident metabolic syndrome, which, in part, is defined by hyperglycemia and treatment for diabetes (11). In the Multi-Ethnic Study of Atherosclerosis, higher diet soda consumption was associated with incident diabetes, incident metabolic syndrome, and individual components of metabolic syndrome (waist circumference and glucose) (12). Others have shown that artificial sweeteners are associated with glucose intolerance through alterations.

Figure 1. | Risk of incident ESRD for highest versus lowest consumption frequency category for diet soda in subgroups of the study population. Adjusted for age, sex, race-center, education level, smoking status, physical activity, total caloric intake, baseline eGFR (linear spline terms with one knot at 90 ml/min per 1.73 m²), body mass index category, diabetes, systolic BP, and serum uric acid. 95% CI, 95% confidence interval; HR, hazard ratio.
### Table 3. Risk (95% CI) of incident ESRD by frequency of sugar-sweetened beverage consumption

<table>
<thead>
<tr>
<th>Categories of SSB Consumption, glasses/wk</th>
<th>Variables</th>
<th>P value for trend</th>
<th>Model 1</th>
<th>Model 2</th>
<th>Model 3a</th>
<th>Model 3b</th>
<th>Model 3c</th>
</tr>
</thead>
<tbody>
<tr>
<td>Continuous SSB Consumption</td>
<td></td>
<td></td>
<td>0.23</td>
<td>0.001</td>
<td>0.05</td>
<td>0.01</td>
<td>0.11</td>
</tr>
<tr>
<td></td>
<td>1 (Ref)</td>
<td>0.90 (0.68 to 1.23)</td>
<td>0.62 (0.42 to 0.90)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
</tr>
<tr>
<td></td>
<td>2 (Ref)</td>
<td>0.90 (0.68 to 1.23)</td>
<td>0.62 (0.42 to 0.90)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
</tr>
<tr>
<td></td>
<td>3 (Ref)</td>
<td>0.90 (0.68 to 1.23)</td>
<td>0.62 (0.42 to 0.90)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
</tr>
<tr>
<td></td>
<td>4 (Ref)</td>
<td>0.90 (0.68 to 1.23)</td>
<td>0.62 (0.42 to 0.90)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
</tr>
<tr>
<td></td>
<td>5–7 (n=5870)</td>
<td>0.90 (0.68 to 1.23)</td>
<td>0.62 (0.42 to 0.90)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
<td>0.81 (0.60 to 1.09)</td>
<td>0.63 (0.42 to 0.92)</td>
</tr>
</tbody>
</table>

**Model 1:** Adjusted for age, sex, race-center, education level, smoking status, physical activity, total caloric intake, baseline eGFR (linear spline terms with one knot at 90 ml/min per 1.73 m²), and Alternative Healthy Eating Index 2010), dietary sodium, dietary fructose, frequency of consumption of diet soda. **Model 2:** Model 1 + dietary phosphorus. **Model 3a:** Model 2 + dietary phosphorus, HR, hazard ratio; 95% CI, 95% confidence interval; SBB, sugar-sweetened beverage; HR, hazard ratio; not applicable.

There are strengths and limitations of our study to acknowledge. The main study limitation is that the dietary data were self-reported and, as such, could be affected by measurement error. Further, although the FFQ can be used to rank individuals according to frequency of consumption of food items (e.g., diet soda), it is not the ideal instrument for quantifying absolute amounts of micronutrients (e.g., dietary intake of phosphorus). Those individuals with diagnosed disease, including diabetes, may have modified their diet for disease management. Specifically, the study participants who were classified as high consumers of diet soda in our study could have substituted this beverage after receiving counseling to reduce their dietary intake of sugar-sweetened beverages for the purpose of weight loss or glycemic control (42). Therefore, dietary intake assessed at visits 1 and 3 may not represent the relevant exposure period for individuals with comorbidities – diabetics and obesity in particular. Another limitation is the lack of measurement of albuminuria. As a result, we were unable to adjust for this covariate and we were unable to incorporate albuminuria into our outcome definition. However, diet soda consumption was not associated with albuminuria in NHANES; thus, albuminuria is not likely to confound the observed association (16). The regression models that include potential mediating factors (models 2–3c) remove some of the true association between diet soda consumption and ESRD risk; thus, these hazard ratios may be underestimated. The association between diet soda and earlier stages of kidney disease was weaker than that for more advanced kidney disease. Residual confounding due to unmeasured or imprecisely measured confounders could, in part, explain the observed association. The primary study strengths mainly relate to the study design. Given the large sample size (n=15,368) and long-term follow-up (median=23 years) in the ARIC study, we were able to ascertain a sufficient number of cases of the clinically-relevant and validated outcome of ESRD (n=357) (21). The ARIC study population is more diverse than prior studies on this topic and is representative of middle-aged black and white men and women from several United States communities.

In conclusion, there was a dose-response relationship between diet soda consumption and ESRD risk. Given the high prevalence of diet soda consumption in the United States, this finding could have a significant public health effect. Further research is needed to validate these findings in other study populations as well as to examine mechanisms through which diet soda could affect kidney disease risk.
Acknowledgments

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Disclosures

None.

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